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SENSITIVITY ANALYSIS AND RELATED ANALYSES: A REVIEW OF SOME STATISTICAL TECHNIQUES

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This paper reviews five related types of analysis, namely (i) sensitivity or what-if analysis, (ii) uncertainty or risk analysis, (iii) screening, (iv) validation, and (v) optimization. The main questions are: when should which type of analysis be applied; which statistical techniques may then be used? This paper claims that the proper sequence to follow in the evaluation of simulation models is as follows. 1) Validation, in which the availability of data on the real system determines which type of statistical technique to use for validation. 2) Screening: in the simulation's pilot phase the really important inputs can be identified through a novel technique, called sequential bifurcation, which uses aggregation and sequential experimentation. 3) Sensitivity analysis: the really important inputs should be subjected to a more detailed analysis, which includes interactions between these inputs; relevant statistical techniques are design of experiments (DOE) and regression analysis. 4) Uncertainty analysis: the important environmental inputs may have values that are not precisely known, so the uncertainties of the model outputs that result from the uncertainties in these model inputs should be quantified; relevant techniques are the Monte Carlo method and Latin hypercube sampling. 5) Optimization: the policy variables should be controlled; a relevant technique is Response Surface Methodology (RSM), which combines DOE, regression analysis, and steepest-ascent hill-climbing. The recommended sequence implies that sensitivity analysis procede uncertainty analysis. Several case studies for each phase are briefly discussed in this paper.

Keywords: Sensitivity analysis; what-if; uncertainty analysis; risk analysis; validation; optimization; regression analysis; least squares; design of experiments; screening; Latin hypercube sampling; perturbation

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1. INTRODUCTION

The *objective* of this paper is to *review the state of the art* in five related types of analysis, namely (i) sensitivity analysis or SA, (ii) uncertainty analysis or UA, (iii) screening, (iv) validation, and (v) optimization. The main questions are: when should which type of analysis be applied; which statistical techniques may then be used? What, however, is meant by these five terms? Since the vast literature on simulation does not provide standard definitions, this paper will use its own descriptions of these terms.

The *main conclusion* will be that the simulation analysts should distinguish the following *five phases* in their simulation studies, and consider the following statistical *techniques* for use in these phases.

Phase 1: Validation

Validation concerns the question: is the simulation model an adequate representation of the corresponding system in the real world? Model validity should be of major interest to both users and analysts. The analysts may use regression and design of experiments or DOE, especially if there are no data on the input/output of the simulation model or its modules. If there are ample data, then the simulation model may be validated by a special type of regression analysis (see §4).

Phase 2: Screening

Screening is needed whenever a simulation study is still in its early, *pilot* phase and *many factors* may be conceivably important (dominant). Straightforward numerical experiments with such a simulation model may require too much computer time. Therefore analysts may use sequential bifurcation, which is a simple, efficient, and effective technique (see §2).

Phase 3: Sensitivity Analysis

This paper uses the term SA (or what-if analysis) for the systematic investigation of the reaction of the simulation response to either *extreme values* of the model's quantitative factors (parameters and input

variables) or to *drastic changes* in the model's qualitative factors (modules). For example, what happens to the mean waiting time in a supermarket, when the arrival rate doubles; what happens if a 'fast lane' is opened for customers with less than (say) seven articles? So this paper's focus is not on marginal changes in inputs (such small perturbations or local sensitivities are discussed in Helton 1993, Ho and Cao 1991, Kleijnen and Rubinstein 1996, and Rubinstein and Shapiro 1993).

If the model's input/output (I/O) transformation is non-monotonic, then it may be dangerous to consider extreme input values only: the output might happen to be approximately the same at the two extremes. Non-monotonicity may be quantified by quadratic effects. Note that this paper concentrates on a single response per run; multiple responses are only briefly mentioned in §6 (on RSM); also see McKay (1995).

In SA the analysts may use *regression analysis* to generalize the results of the simulation experiment, since regression analysis characterizes the I/O transformation of the simulation model. To obtain better estimators of the effects in the regression model, these estimators should be based on DOE. DOE requires fewer simulation runs than intuitive designs that change one factor at a time ('ceteris paribus' approach). See §3.

Phase 4: Uncertainty Analysis

In UA (also called risk analysis in some publications) the values of the model inputs are sampled from prespecified distributions, to quantify the consequences of the uncertainties in the model inputs on the model outputs. So in UA the input values range between the extreme values investigated in SA. The goal of UA is to quantify the *probability* of specific output values, whereas SA does not tell how likely a specific result is. The differences between SA and UA will be further explored later on (also see the discussion in Draper 1995).

In UA the analysts use the technique of *Monte Carlo* sampling, possibly including variance reduction techniques such as *Latin hypercube sampling* or LHS, possibly combined with regression analysis (see §5).

Phase 5: Optimization

From the users' viewpoint the important model inputs should be split into two groups, namely inputs that are under the users' *control*

(policy variables) versus *environmental* inputs, which are not controllable. Specifically, users want to ask 'what if' questions: what happens if controllable inputs are changed (scenario analysis), what if other inputs change (UA)? To *optimize* the controllable inputs, the analysts may use *Response Surface Methodology* or RSM, which builds on regression analysis and DOE (see §6). Note that in nuclear engineering the term RSM refers to SA without optimization; see Olivi (1980) and Rao and Sarkar (1995).

This list of five phases implies that SA should precede UA. McKay (1995, pp. 7, 33) also discusses the sequence of steps in modelling, but he concentrates on UA.

SA is applied in simulation studies of very different real-life systems, in all kinds of disciplines: chemistry, physics, engineering, economics, management science, and so on. Moreover, the theoretical aspects of SA are studied in mathematics and statistics. Unfortunately, the definition of SA (and of related analyses) varies over and within these many disciplines! Because terminologies differ so much, communication among different disciplines is difficult. Yet, *cross-fertilization* among these disciplines is certainly possible and fruitful.

Especially the roles of SA and UA seem unclear (see Kleijnen 1994). It seems wise to consider questions that might be asked by the *users* of simulation models (these users are the clients of the simulation modelers, and 'the customer is king'). Examples are: (i) What is the probability of a nuclear accident happening at out site? (ii) How big is the chance that the financial investment in our project will turn out to be unprofitable? (iii) What is the probability of customers having to wait longer than three minutes at our supermarket's checkout lanes?

Next consider examples of questions asked by *simulationists* (or analysts): (i) What happens if parameter h (with $h = 1, \ldots, k$) of our simulation model is changed? (ii) Which are the really important parameters among the hundreds of parameters of our total simulation model (that consists of several modules or submodels)?

Uncertainty about a simulated system's response (outcome) may have two different causes:

(i) The system's process is well-known, so it is represented by a *deterministic* model; however, its *parameters*, are not known exactly.

Many examples can be found in classical physics. Karplus (1983) gives an interesting survey of different types of models in different disciplines, ranging from astronomy through ecology to social sciences; these disciplines use white box, grey box, and black box models respectively.

(ii) Some models are *intrinsically stochastic*: without the randomness the problem disappears. In the supermarket example the customer arrival and service times are random, which creates stochastic customer waiting times. In an ecological model an example may be the wind's speed and direction. Helton (1993, 1996) speaks of 'stochastic or aleatory uncertainty'. Also see Van Asselt and Rotmans (1995, pp. 11–13) and Zeigler (1976).

Uncertainty is the central problem in *mathematical probability theory*. This discipline, however, has two schools (also see Cooke 1995, pp. 4–6, and Winkler 1996):

(i) The *objectivists or frequentists* school: for example, when throwing a dice many times, the probability of throwing a six can be defined as the limit of the frequency of throwing a six (obviously this limit is 1/6 for a fair dice).

(ii) The *subjectivists* school: an event may be unique (for example, tomorrow's weather at a particular place); yet a 'probability' (between zero and one) may be assigned to a particular outcome (for example, sunny all day). Many UA studies concern such unique events; for example, the following questions by the users were mentioned above: (i) What is the probability of a nuclear accident happening at our site? (ii) How big is the chance that our financial investment in this project will turn out to be unprofitable? Note that *Bayesians* try to combine prior subjective data with new factual data; see Draper (1995).

Zeigler wrote a seminal book on the theory of modeling and simulation, in which he distinguishes between *input variables* and *parameters*; see Zeigler (1976). By definition, a variable can be directly observed; an example is the number of check-out lanes in a supermarket. A parameter, however, can *not* be observed directly, so its value must be inferred from observations; an example is the arrival rate of customers. Hence mathematical statistics may be used to quantify the probability of certain values for a parameter (see §5.1). When applying

DOE (for example, a fractional factorial design) to simulation models, this paper will use the term *factor* for parameters, input variables, and modules that are changed from (simulation) run to run. (For modules treated as factors also see Helton *et al.*, 1995a, Mckay 1995, pp. 51–54, and Van Asselt and Rotmans 1995.)

This paper covers both practical and theoretical aspects. It covers these aspects at an advanced tutorial level; that is, the readers are assumed to be experienced professionals in simulation. For technical details, a bibliography with approximately 80 references is included. (To reduce the number of references, only the most recent reference for a certain technique or concept is given; this recent reference may refer to the original, older publication.) The review synthesizes the latest research results in a unified treatment of the various analysis types. It also briefly discusses selected applications of these analyses. The statistical techniques mentioned above, have already been applied many times in practical simulation studies, in many domains (also see the many contributions in this special issue). These techniques make simulation studies give more general results, in less time. Unfortunately, nobody can be an expert in the many disciplines that discuss the issues addressed in this paper! This paper is *colored* by more than 25 years of experience with simulation, especially its statistical aspects and its applications to problems in business, economics, environmental, agricultural, military, and computer systems. This paper combines, updates, and revises Kleijnen (1994, 1995b).

For didactic reasons, this paper is not organized in the order of the five stages given above; for example, validation (phase 1) may use DOE, but DOE is simpler explained for SA (stage 3). So in §2 screening and especially sequential bifurcation are discussed. In §3 SA through regression analysis and DOE is explained. In §4 validation is briefly addressed. In §5 UA is discussed, distinguishing between deterministic and stochastic simulations. In §6 optimization through RSM is explained. In §7 conclusions are summarized. A bibliography ends this paper.

2. SCREENING: SEQUENTIAL BIFURCATION

Screening is the search for the few (say) k_1 really important factors among the great many (say) k_2 potentially important factors ($k_1 \ll k_2$).

In practice, experiments with simulated systems often do involve many factors. Andres (1996) fives an example with 3300 factors that affect deep geological waste disposal. Rahni *et al.* (1996) screen 390 factors in a deterministic simulation of an experimental building, developed to study energy flows in buildings. Another example is the case study in this section, which has 281 factors; this study concerns a complicated deterministic simulation model of the 'greenhouse' phenomenon (the increase in temperatures worldwide, caused by increasing quantities of carbon dioxide or CO_2 and of other gases in the atmosphere). McKay (1995) discusses case studies with 84 and 36 factors respectively. Other examples would be stochastic simulations of queueing networks, such as logistic, computer, and telecommunication networks with nodes that represent individual servers (machines). These queueing models have many parameters (service and arrival rates), input variables (number of servers), and behavioral relationships (queuing discipline or priority rule). In summary, simulation models with *several hundred factors* are common; also see Morris (1991).

The problem is that a simulation run may require so much *computer time* that the number of runs (say) n must be much smaller than the number of factors ($n \ll k_2$). For example, the greenhouse model with 281 factors takes 15 minutes of computer time per run, so 282 runs would have taken roughly 70 hours of computer time (on the computer available at that time: PC 386DX, 16 Mhz). In general, the analysts and their clients may not wish to wait until (roughly) k_2 runs will have been generated. (The greenhouse model is part of an even bigger model, called 'IMAGE', Integrated Model to Assess the Greenhouse Effect, which was developed at the Dutch institute RIVM, National Institute of Public Health and Environmental Protection; see Rotmans 1990.)

Whenever there are many factors the analysts should assume that only a few factors are really important ($k_1 \ll k_2$): *parsimony, Occam's razor*. Moreover, to solve the screening problem ($k_1 \ll k_2$, $n \ll k_2$), one general scientific principle can be applied: *aggregation*. Indeed, aggregation has been used in the study of many large systems; for example, in economic theory and practice the production volumes of the individual companies are aggregated into the Gross National Product or GNP.

More specifically, in the experimentation with large systems (real or simulated), analysts have applied *group screening*: they combine individual factors into groups and experiment with these groups as if they were individual factors. The theory on group screening goes back a long time: Jacoby and Harrison (1962), Li (1962), Patel (1962), and Watson (1961). A more recent publication is Morris (1987).

Practical applications of group screening, however, are rare. The reason may be that in experiments with real systems it is impossible to control hundreds of factors. And in simulation, most analysts seem to be unaware of group screening. Four simulation applications of classical group screening are summarized in Kleijnen (1987, pp. 327); other applications are given in Cochran and Chang (1990) and Rahni *et al.* (1996).

This paper focusses on a *novel* group screening technique, namely *sequential bifurcation* or SB. SB uses a special design; it also applies a different analysis. SB proceeds sequentially (or stagewise) and splits up (or bifurcates) the aggregated inputs as the experiment proceeds, until finally the important individual inputs are identified and their effects are estimated. SB is more efficient than competing group screening techniques. Moreover, SB has been used in practice, namely in the greenhouse simulation mentioned above, and in the building thermal deterministic simulation in De Wit (1995). For the greenhouse simulation, SB found the 15 most important inputs among the 281 factors after only 144 runs. SB lead to surprising results: the technique identified some factors as being important that the ecological experts assumed to be unimportant. In De Wit's building simulation, SB found the 16 most important inputs among the 82 factors after only 50 runs. De Wit checked these results by applying a different screening technique, namely 'randomized one-factor-at-a time designs' of Morris (1991), which took 328 runs.

How general are the results of these SB applications; can SB be applied with confidence to other simulation models? Scientists must always make *assumptions* to make progress; more specifically, any screening technique must use assumptions. In practice, the simulation analysts often leave their assumptions implicit. Frequently the analysts assume that they know which factors are unimportant, and they investigate only a few *intuitively* selected factors; also see Bankes and Lucas (1995). (Often they apply an inefficient or ineffective design: they

change one factor at a time; see Van Groenendaal (1994). All classical group screening techniques and also SB assume a low-order polynomial approximation or metamodel for the I/O transformation of the simulation model; moreover they assume known signs or directions for the first-order or main effects. These assumptions will be discussed next.

Assumption 1: Low-Order Polynomial Approximation

A metamodel implies that the underlying simulation model is treated as a *black box*. The advantage of a low-order polynomial is that it is simple and that it may apply to *all* types of random and deterministic simulation. The disadvantage is that the approach cannot exploit the special structure of the simulation model at hand. Note that other techniques do use the specific structure of the specific simulation model; examples are importance sampling or score function (see §5.1) and differential or perturbation analysis (see Helton 1993 for continuous systems and Ho and Cao 1991 for discrete-event systems). However, these sophisticated techniques have not yet been applied to discrete-event simulations with hundreds of factors.

Low-order polynomials are often used in DOE with its concomitant Analysis of Variance (ANOVA), applied to real or simulated systems with a small number of factors; for a survey see Kleijnen and Van Groenendaal(1992).

All classic group screening designs assume a first-order polynomial. In SB a first-order polynomial requires only half the number of runs that an approximation with two-factor interactions does (foldover principle; see §3.3.2 and also Andres 1996 and Kleijnen 1987). In general, however, a more cautious approach is recommended, namely a metamodel with interactions.

Assumption 2: Known Signs

In group screening the analysts must assume known signs or directions for the effects, in order to know with certainty that individual effects do *not compensate* each other within a group. In practice the sign of a factor may be known indeed (the magnitude, however, is unknown, so simulation is used). For example, in the greenhouse study the ecological experts felt confident when they had to specify the signs of the factor effects, for most factors. In queueing networks the

response may be throughput, so the analysts may assume that higher speeds of servers have non-negative effects on the response; and so does an increasing arrival rate of jobs at the source of the network.

If the analysts feel that they do *not* know the signs of *a few* factors, then they may treat these factors separately. Indeed, in the greenhouse case study there is a very small group of factors with unknown signs. These factors can be investigated in a traditional design.

A special characteristic of sequential procedures such as SB is that the analysts do not need to quantify *a priori* how big a factor effect should be in order to be called *important*. As simulation outputs are generated sequentially, SB computes upper limits for the factor effects, and the analysts can stop the simulation experiment as soon as they find these limits sharp enough. Obviously, the analysts may make these limits depend on the system being simulated.

The main objective of this section was to inform simulation analysts about a novel technique for the screening of large simulation models. Therefore concepts were emphasized; for technical details the readers are referred to the more than 200 pages of the doctoral dissertation, Bettonvil (1990) or to the summary paper by Bettonvil and Kleijnen (1996). Different screening techniques are Andres's (1996) Iterated Fractional Factorial Design (IFFD) and McKay's (1995) replicated LHS design.

3. SENSITIVITY ANALYSIS: REGRESSION ANALYSIS AND DOE

Sensitivity analysis was defined (in §1) as the systematic investigation of the reaction of the simulation response to either *extreme* values of the model's quantitative factors (parameters and input variables) or to drastic changes in the model's qualitative factors (modules). This section assumes that the number of factors is relatively small, that is, the screening phase is over. An example with fourteen factors will be mentioned in §6 (section on optimization). In the sensitivity phase, regression analysis may be used to approximate the I/O transformation of the simulation model. This regression analysis gives better results when the simulation experiment is well designed, using classical statistical designs such as fractional factorials.

3.1. Introduction: Graphical Methods

Practitioners often make a *scatter plot* with on the y-axis the simulation response (say, average waiting time) and on the x-axis the values of one factor (for example, service rate). This graph indicates the I/O transformation of the simulation model, treated as a black box. It shows whether this factor has a positive or negative effect on the response, whether that effect remains constant over the domain or experimental area of the factor, and so on. Also see Helton (1993, pp. 347–349).

The practitioners may further analyze this scatter plot: they may fit a curve to these (x, y) data, for example, a straight line ($y = \beta_0 + \beta_1 x$). Of course, they may fit other curves (such as a quadratic curve: second degree polynomial) or they may use paper with one or both scales logarithmic.

To study *interactions* between factors, they may combine scatter plots per factor. For example, they drew one scatter plot for different values of the service rate, given a certain number of servers. They can now superimpose plots for different numbers of servers. Intuitively, the waiting time curve for a low number of servers should lie above the curve for a high number of servers (if not, the simulation model is probably wrong; see the discussion on validation in §4). If the response curves are not parallel, there are interactions, by definition.

Making scatter plots is a simple technique that is often effective. However, superimposing many plots is cumbersome. Moreover, their interpretation is subjective: are the response curves really parallel and straight lines? These shortcomings are removed by regression analysis.

3.2. Regression Analysis

A regression metamodel was defined as an approximation of the I/O transformation of the underlying simulation model. Consider the *second degree polynomial*

$$Y_i = \beta_0 + \sum_{h=1}^{k} \beta_h x_{i,h} + \sum_{h=1}^{k} \sum_{h'=h}^{k} \beta_{h,h'} x_{i,h} x_{i,h'} + E_i \qquad (1)$$

$$(i = 1, \ldots, n)$$

with

Y_i : simulation response of factor combination i (stochastic variables are shown in capitals);
β_0 : overall mean response or regression intercept;
β_h : main or first-order effect of factor h;
$x_{i,h}$: value of the standardized factor h in combination i (see Equation (2) below);
$\beta_{h,h'}$: interaction effect of the factors h and h' with $h<h'$;
$\beta_{h,h}$: quadratic effect of factor h;
E_i : fitting error of the regression model for factor combination i;
n : number of simulated factor combinations.

For didactic reasons, interactions and quadratic effects are ignored initially. Then the *relative importance* of a factor is obtained by sorting the absolute values of the main effects β_h, provided the factors are *standardized.* So let the original (non-standardized) factor h be denoted by z_h. In the simulation experiment z_h ranges between a lowest value (say) l_h and an upper value u_h; that is, either the simulation model is not valid outside that range (see the discussion on validation in §4) or in practice that factor can range over that domain only (for example, the number of servers can vary only between one and five). (DOE concerns the question whether z_h is to be set at the extreme values only or also at intermediate values; see §3.3.) The variation (or spread) of factor h is measured by $a_h=(u_h-l_h)/2$; its location (or mean) by $b_h=(u_h+l_h)/2$. Then the following standardization is appropriate:

$$x_{ih}=(z_{ih}-b_h)/a_h. \tag{2}$$

Note that other measured besides $|\hat{\beta}_h|$ have been proposed; see McKay (1995); also see Helton (1993), Saltelli (1996), and Sobol (1995). More research and applications are needed.

The classic *fitting algorithm*, which determines the regression parameters $\boldsymbol{\beta}=(\beta_0,\beta_1,\ldots,\beta_{k,k})'$ in Equation (1), uses the *ordinary least squares* (OLS) criterion. Software for this algorithm is abundant. If, however, *statistical assumptions* about the fitting error are added, then there are better algorithms. Consider the following assumptions.

It is realistic to assume that the variance of the stochastic fitting error E varies with the input combination of the random simulation model: $\text{var}(E_i) = \sigma_i^2$. (So both the mean and the variance of Y, the response of the stochastic simulation, depend on the input.) Then *weighted least squares* or WLS (with the standard deviations σ_i as weights) yields unbiased estimators of the factor effects, but with smaller variances than OLS gives.

To improve the SA, *common* pseudorandom number seeds should be used when simulating different factor combinations for a random simulation model. Then, however, *generalized least squares* or GLS should be applied to get minimum variance, unbiased estimators. Unfortunately, the variances and covariances of the simulation responses Y_i are unknown, so they must be estimated. The following equation gives the classic covariance estimator, assuming d independent replications or runs per factor combination (so Y_{ig} and $Y_{i'g}$ are correlated, but Y_{ig} and $Y_{ig'}$ are not):

$$\hat{\text{cov}}(Y_i, Y_{i'}) = \sum_{g=1}^{d} (Y_{ig} - \bar{Y}_i)(Y_{i'g} - \bar{Y}_{i'})/(d-1). \tag{3}$$

Under mild conditions, the resulting *estimated* GLS gives unbiased estimators of the factor effects, but with smaller variances than OLS gives; see Kleijnen and Van Groenendaal (1992).

To make statistical inferences (for example, about the importance of the factor effects), a *Gaussian* distribution is normally assumed. To satisfy this assumption, the analysts may apply transformations such as the logarithmic transformation fo the simulation response Y. Alternatively, the analysts might hope that the results are not sensitive to 'mild' non-normality. Another alternative is to replace both the simulation response Y and the factors x_h in Equation (1) by their ranks: rank regression. Moreover, the analysts may examine the original and the transformed data, and see whether the various analyses give the same qualitative results. See Andres (1996), Kleijnen (1987), Kleijnen, Bettonvil, and Van Groenendaal (1995), and Saltelli (1996).

Of course, it is necessary to check the fitted regression metamodel: is it an adequate approximation of the underlying simulation model? Therefore the metamodel may be used to predict the outcome for a *new* factor combination of the simulation model. So replace β in the

specified metamodel by its estimate $\hat{\beta}$, and substitute a new input combination (there are n old combinations). Compare the regression prediction $\hat{y}_{n+1}$ with the simulation response y_{n+1}.

A refinement is *cross-validation*: do not add a new combination (which requires computer time), but eliminate one old combination (say combination i) and re-estimate the regression model from the remaining $n-1$ combinations. Repeat this elimination for all values of i (with $i=1,...,n$; see Equation (1)). Statistical details are discussed in Kleijnen and Van Groenendaal (1992); also see Helton (1993, pp. 347–356).

Applications of regression metamodeling will be discussed below. Note that the analysts may use a covariance stationary process (instead of white noise) to model the systematic effects of the inputs; see Sacks, Welch, Mitchell, and Wynn (1989).

3.3. Design of Experiments

The preceding subsection (§3.2) used regression metamodels to approximate the I/O transformation of simulation models. Such a metamodel has (say) q regression parameters in the vector $\boldsymbol{\beta}$, which measure the effects of the k factors; for example, q equals $k+1$ if there are no high-order effects (see §3.3.2).

It is obvious that to get unique, unbiased estimators of these q effects, it is necessary to simulate at least $n \geq q$ factor combinations. Moreover, which n combinations to simulate, can be determined such that the accuracy of the estimated factor effects is maximized (variance minimized). This is the goal of the statistical theory on DOE (which Fisher started in the 1930s, and Taguchi continues today).

3.3.1. Main Effects Only

Consider a first-order polynomial, which is a model with only k main effects, besides the overall mean (see the first two terms in the right-hand side of Equation (1)). By definition, a *resolution III* or R-3 design permits the unbiased estimation of such a first-order polynomial. The following types of R-3 designs are of special interest.

In *practice*, analysts usually first simulate the 'base' situation, and next they change *one factor at a time*; hence $n=1+k$. See Helton *et al.* (1995a, pp. 290) and Van Groenendaal (1994).

Statistical theory on DOE, however, concentrates on *orthogonal designs*, that is, designs that satisfy

$$x'x = nI \tag{4}$$

where **bold** letters denote matrices; $x = (x_{ij})$ denotes the so-called design matrix with $i = 1, \ldots, n$ (so i is defined as in Equation (1)) and $j = 0, 1, \ldots, k$ (obviously, $n > k$; see the beginning of §3.3); the first column of x has the values $x_{i0} = 1$ ('dummy' factor) and corresponds with β_0; the remaining columns have the values x_{ih} defined below Equation (1); finally, I denotes the identity matrix (so this capital letter does not denote a stochastic variable). Orthogonal designs give estimators of $\boldsymbol{\beta}$ that are unbiased and have *smaller variances* than the estimators resulting from designs that change one factor at a time.

Orthogonal designs are tabulated in many publications. The analysts may also learn how to construct those designs; see Kleijnen (1987). Recently, software has been developed to help the analysts specify these designs; see Ören (1993) and Vollebregt (1996).

A well-known class of orthogonal designs are 2^{k-p} fractional factorials; for example, a simulation with $k = 7$ factors requires $n = 2^{7-4} = 8$ factor combinations (runs) to estimate the main effects plus overall mean. Actually, these 2^{k-p} designs also require 8 runs when $4 \leqslant k \leqslant 7$. See Kleijnen (1987).

References to many *simulation applications* of these designs can be found in Kleijnen (1987) and Kleijnen and Van Groenendaal (1992).

In practice, however, it is unknown whether only main effects are important. Therefore orthogonal designs with $n \approx k + 1$ should be used only in optimization (see §6). Moreover, these designs are useful as building blocks if interactions are accounted for; see §3.3.2

3.3.2. Main Effects Biased by Interactions?

It seems prudent to assume that *interactions* between pairs of factors (two-factor interactions) may be important. By definition, a *resolution IV* or R-4 design permits the unbiased estimation of all k main effects, even if two-factor interactions are present; R-4 designs do not give unbiased estimators of all $k(k-1)/2$ individual two-factor interactions. Compared with R-3 designs, R-4 designs require that the number of simulated factor combinations be *doubled*; for example, $k = 7$ now

requires $n = 2 \times 8 = 16$ runs (a R-3 design requires only $n = 2^{7-4} = 8$; see §3.3.1). To construct an R-4 design, the design matrix x of the R-3 design is augmented with $-x$ ('mirror' image; 'foldover' principle; also see §2). R-4 designs may give an indication of the importance of (confounded, aliased, biased) interactions (also see §3.3). Obviously, designs that change one factor at a time (see §3.3.1) do not enable estimation of interactions!

Details, including *simulation applications* are presented in Kleijnen (1987) and Kleijnen and Van Groenendaal (1992). Recent applications include the simulation of a decision support system (DSS) for the investment analysis of gas pipes in Indonesia, and a simulation model for the Amsterdam police; see Van Groenendaal (1994) and Van Meel (1994) respectively.

3.3.3. Individual Factor Interactions

Suppose the analysts wish to estimate the *individual two-factor interactions* $\beta_{h,h'}$ with $h < h'$; see Equation (1). There are $k(k-1)/2$ such interactions (see §3.3.2). By definition, a *resolution V* or R-5 design permits the unbiased estimation of all k main effects and all $k(k-1)/2$ two-factor interactions. Obviously many more simulation runs are now necessary; for example, $k = 8$ factors implies $q = 37$ effects (namely, $1 + 8 + 8 \times 7/2$); another example is $k = 11$, which implies $q = 67$ (namely, $1 + 11 + 11 \times 10/2$). Therefore practitioners study only small values of k; for example, $k < 5$ requires *full factorial two-level designs*, denoted as 2^k. For higher values of k, however, *fractional* factorials are recommended, for example, $2^{8-2}(n = 64)$ for $k = 8$ (so $q = 37$), and $2^{11-4}(n = 128)$ for $k = 11$ (so $q = 67$). Kleijnen (1987) gives details, including construction of R-5 designs, and applications.

If all 2^k combinations are simulated, then high-order interactions (not only two-factor interactions) can be estimated. However, these interactions are hard to interpret: what does $\beta_{1,2,3,4}$ mean? Hence, either transformations such as the logarithmic transformation may be applied or the experimental domain may be restricted. An example of a full factorial with $k = 6$ is provided in Rao and Sarkar (1995). Also see Kleijnen (1987).

3.3.4. Quadratic Effects: Curvature

If all k factors are quantitative and their k quadratic effects $\beta_{h,h}$ ($h = 1, \ldots, k$) in Equation (1) are to be estimated, then k extra runs

may be needed. Moreover, each factor must be simulated for more than two values (taking the extreme values minimizes the variance of the estimated main effects; see §3.3.1).

Popular in statistics and in simulation are *central composite designs*. They consists of a R-5 design (to estimate the intercept β_0, the main effects, and the two-factor interactions; see §3.3.3), $2k$ axial points (namely $-c$ and $+c$ along each of the k axis, with $c \neq 1$ in the standardized factors x_h), and the central point (the origin in the standardized factors). Hence, these designs have five values per factor. Consider an example with $k = 2$ factors. Then there are $q = 6$ effects. But $n = 9 (= 2^2 + 2 \times 2 + 1)$ factor combinations are simulated. In general, these designs require many runs: $n >> q$. An application in nuclear engineering is given by Rao and Sarkar (1995). For more details see Kleijnen (1987) and Kleijnen and Van Groenendaal (1992).

Applications are found in the *optimization* of simulation models (see §6). Note that simulating as many as five values per factor resembles UA, in the sense that the range of factor values is well covered. The metamodeling methodology of this section(3) is discussed at length in Kleijnen and Sargent (1996).

4. VALIDATION

This paper confines the discussion of validation to the role of SA(§3) in validation; other statistical techniques for validation and verification are discussed in Kleijnen (1995a), (Validation of knowledge-based systems is discussed in Cardenosa 1995.) Obviously, validation is one of the first problems that must be solved in a simulation study (stage 1); for didactic reasons, however, validation is discussed now (§4).

True validation requires that *data* on the real system be available. In practice, the amount of data varies greatly: data on failures of nuclear installations are rare, whereas data on electronically monitored systems (such as computers and supermarkets) are abundant.

If data are available, then many statistical techniques can be applied. For example, simulated and real data on the response, can be compared through the Student statistic for paired observations, assuming the simulation is fed with real-life input data: *trace driven simulation*. A better test uses well-known regression analysis, but does

not test whether real and simulated data lie on a straight line through the origin! Instead, the difference between simulated and real data is regressed on their sum; this novel regression test is discussed in Kleijnen, Bettonvil, and Van Groenendaal (1996).

However, if no data are available, then the following type of SA can be used. The analysts and their clients do have *qualitative* knowledge of certain parts of the real system; that is, they do know in which direction certain factors affect the response of the corresponding module in the simulation model (also see the discussion on sequential bifurcation in §2). If the regression metamodel (see §3) gives an estimated factor effect with the *wrong sign*, this is a strong indication of a wrong simulation model or a wrong computer program.

Applications are given in Kleijnen, Van Ham, and Rotmans (1992), who discuss the greenhouse model IMAGE, and Kleijnen (1995c), who discusses a military model, namely the hunt for mines on the bottom of the sea. These applications further show that the validity of a simulation model is restricted to a certain domain of factor combinations. This domain corresponds with the *experimental frame* in Ziegler (1976), defined as the limited set of circumstances under which the real system is to be observed or experimented with.

Moreover, the regression metamodel shows which factors are most important. For the important *environmental inputs* the analysts should try to collect data on the values that occur in practice. If they do not succeed in getting accurate information, then they may use the UA of the next section.

5. UNCERTAINTY ANALYSIS: MONTE CARLO AND LATIN HYPERCUBE SAMPLING

As the preceding section mentioned, the analysts may be unable to collect reliable data on important environmental inputs; that is, the values that may occur in practice are uncertain. Then the analysts may apply UA. The goal of UA is to quantify the *probability* of specific output values, whereas SA (as defined in §1) does not tell how likely a specific result is. The differences between SA and UA are further explored below.

5.2. The Basics of Uncertainty Analysis

First the analysts derive a *probability function* for the input values. This distribution may be estimated from sample data, if those data are available; otherwise this distribution must be based on subjective expert opinions (also see Draper 1995, pp. 92, Helton 1993, pp. 337–341, Helton *et al.*, 1992, chapter 2, pp. 4, Helton *et al.*, 1995a, pp. 288, Kraan and Cooke 1995, Moors *et al.* 1995). Popular distribution types are uniform, loguniform, triangular, beta, normal, and lognormal distributions. Usually the inputs are assumed to be statistically independent. Nevertheless, correlated inputs are discussed in Bedford and Meeuwissen (1996), Helton (1993, pp. 343–345), Cooke (1995), Helton *et al.* (1992, chapter 3, pp. 7), Reilly, Edmonds, Gardner, and Brenkert (1987), and Reilly (1996).

Next the analysts use pseudorandom numbers to sample input values from those distributions: *Monte Carlo or distribution sampling.* UA often uses *Latin hypercube sampling* (LHS), which forces the sample of size (say) n to cover the whole experimental area; for example, in case of a single input, this input's domain is partitioned into (say) s equally likely subintervals and each subinterval is sampled s/n times. See Helton (1993, pp. 341–343).

This paper's message (which is certainly bound to be contended) is that LHS is recommended as a *variance reduction technique* or VRT, not as a *screening* technique. For screening technique purposes the inputs should be changed to their extreme values, whereupon their effects should be computed; see the discussion on screening in §2. Of course, the larger sample in LHS gives more insight than the small sample in screening does; however, for a large number of factors such a large sample is assumed to be impossible. Also see Banks (1989) versus Downing *et al.* (1986) and McKay (1992).

The sampled input values are fed into the simulation model. This subsection focuses on *deterministic* simulation models (the next subsection covers stochastic models). Hence, *during* a simulation run all its inputs are deterministic; for example, the input is constant or shows exponential growth. From run to run, however, the (sampled) inputs vary; for example, constants or growth percentages change. These sampled inputs yield an estimated distribution of ouput or response values. That distribution may be characterized by its location (measured by the mean, modus, or median) and its dispersion (quantified by

the standard deviation or various percentiles and quantiles, such as the 90% quantile). For a basic introduction to UA see Helton (1993) and Kleijnen and Van Groenendaal (1992, pp. 75–78).

Which quantities sufficiently summarize a distribution function, depends on the users' *risk attitude*: risk neutral (in that case the mean is a statistic that characterizes the whole distribution sufficiently), risk aversion, or risk seeking; see Balson, Welsh, and Wilson (1992) and Bankes (1993, pp. 444). The former authors further distinguish between *risk assessment* (defined as UA in this paper) and *risk management* (risk attitude, possible countermeasures); also see Brehmer, Eriksson, and Wulff (1994), Hora (1996), and Van Asselt and Rotmans (1995).

Combining UA with *regression analysis* gives estimates of the effects of the various inputs; that is, regression analysis shows which inputs contribute most to the uncertainty in the output. (Mathematically, this means that in Equation (1) the deterministic independent variables $x_{i,h}$ are replaced by random variables $X_{i,h}$). Because more values are sampled per factor, more complicated metamodels might now be used. Indeed, for prediction purposes these metamodels may be made really complicated; for example, splines may be used. For explanatory purposes and SA, however, simple metamodels may be preferred; also see Kleijnen (1979), and Kleijnen and Sargent (1996). Note that Helton *et al.* (1991, 1992) call this combination of uncertainty and regression analysis 'sensitivity analysis'.

UA is applied in *business and economics*. Hertz (1964) introduced this analysis into investment analysis: what is the probability of a negative Net Present Value? Krumm and Rolle (1992) give recent applications in the Du Pont company. Birge and Rosa (1995) and Van Groenendaal and Kleijnen (1996) also discuss investment analysis issues. UA in business applications may be implemented through add-ons (such as @RISK and Crystal Ball) that extend spreadsheet software (such as Lotus 1-2-3 and Excel). Moreover, these add-ons are augemented with distribution-fitting software (such as ExpertFit) and optimization software (such as What's Best).

In the *natural sciences*, UA is also popular. For example, in the USA the Sandia National Laboratories have performed many uncertainty analyses for nuclear waste disposal and reactor safety (Breeding *et al.*, 1992, Helton and Breeding 1993, Helton *et al.*, 1991, 1992, Hora 1996). Oak Ridge National Laboratory has investigated radioactive

doses absorbed by humans (Downing *et al.*, 1985). Nuclear reactor safety has been investigated for the Commission of the European Communities (Olivi 1980, Saltelli and Homma 1992). UA has also been performed at the Dutch RIVM (Harbers 1993, Janssen *et al.*, 1992). Three environmental studies for the electric utility industry were presented in Balson *et al.* (1992). UA in the natural sciences has been implemented through software such as LISA (see Saltelli and Homma 1992, pp. 79), PRISM (Reilly *et al.*, 1987), and UNCSAM (Janssen *et al.*, 1992).

Note that UA is also used in the analysis of computer security; see Engemann and Miller (1992) and FIPS (1979).

The beginning of this section (§5) mentioned that a basic characteristic of UA is that information about the inputs of the simulation model is not reliable; therefore the analysts do not consider a single 'base value' per input variable, but a distribution of possible values. Unfortunately, the form of that distribution must be specified (by the analysts together with their clients). There is the danger of *software driven specification*; that is, the analysts concentrate on the development of software that implements a variety of statistical distributions, but their clients are not familiar at all with the implications of these distributions; also see Easterling (1986). Bridging this gap requires intensive collaboration between model users, model builders, and software developers. Consequently, it may be necessary to study the effects of the *specification* of the input distributions (and of other types of inputs such as scenarios). This type of SA may be called *robustness analysis*. Examples are given by Helton *et al.* (1992, section 4.6) and Helton *et al.* (1995b); also see Helton *et al.* (1995a), Janssen *et al.* (1992), Kleijnen (1987, pp. 144–145), and McKay (1995, pp. 31).

Robustness analysis may also use more sophisticated, faster sampling techniques that are based on *importance sampling or likelihood ratios*, which changes the original input distribution. Technical details can be found in Beckman and McKay (1987), Kleijnen and Rubinstein (1996), and Rubinstein and Shapiro (1993).

Note that importance sampling is also very useful (if not indispensable) whenever *rare events* must be simulated, such as nuclear accidents and buffer overflows in reliable telecommunication networks. See Helton *et al.* (1995a, pp. 290), Kleijnen and Rubinstein (1995), Rubinstein and Shapiro (1993), and Sarkar and Rief (1995).

5.2. Uncertainty Analysis of Stochastic Models

The type of question answered by UA is 'what is the chance of ...?' So the model must contain some random element. In §5.1 that randomness was limited to the inputs of the model, whereas the model itself was deterministic. However, as the Introduction (§1) mentioned, some models are intrinsically *stochastic*: without the randomness the problem disappears. Examples are queueing models, where the customer interarrival times may be independent drawings from an exponential distribution with parameter λ (so it mean is $1/\lambda$). This parameter is an input of the simulation queueing model. That model generates a stochastic time series of customer waiting times. The question may be: what is the probability of customers having to wait longer than 15 minutes? For simple queueing models this question can be answered analytically or numerically, but for more realistic models the analysts use simulation. Mathematical statistics is needed to determine how many customers must be simulated in order to estimate the response with prespecified accuracy; see Kleijnen and Van Groenendaal (1992, pp. 187–197). (Helton *et al.*, 1995a, p. 287 state: 'Stochastic uncertainty is a property of the system being studied, while subjective uncertainty is a property of the analysts performing the study'.)

How to apply UA to such a queueing simulation? Suppose the interarrival parameter λ is estimated from a sample of r independent interarrival times. Then the *central limit theorem* implies that the distribution of the estimated interarrival parameter $\hat{\lambda}$ approximates a normal distribution. Hence the parameter value $\hat{\lambda}^*$ can be sampled from this distribution, and be used as input to the queueing simulation. That simulation is run for 'enough' customers. Next the procedure is repeated: sample $\hat{\lambda}^*$, and so on. For details see Kleijnen (1983).

Instead of relying on the central limit theorem, Cheng and Holland (1996) apply *bootstrapping*. However, the question remains *which* response to report to the users: the unconditional, ex post variance as do Cheng and Holland (1996) and also Brady and Hillestad (1995, p. 30); the ex post variance, mean, and various quantiles, as Haverkort and Meeuwissen (1995) do; or the conditional moments (conditioned on the values of the estimated parameters)? The discussion in Draper

(1995, pp. 78, 83) clearly demonstrates how controversial this issue is. Also see Kleijnem (1983).

In summary, UA has hardly been applied to stochastic models such as queueing models (SA has been employed in many simulation studies; see §3). Helton (1993, pp. 356–358) and Helton *et al.* (1991, 1992, 1995a) discuss UA of stochastic models in the natural sciences (nuclear power plants, the spreading of nuclides). So UA of stochastic simulation models is an interesting area for further research.

6. OPTIMIZATION: RESPONSE SURFACE METHODOLOGY

The decision makers should control the *policy variables*. For example, in the greenhouse case the government should restrict emissions of the gases concerned; in queueing problems, management may add more servers (such as check-out lanes at a supermarket). Strictly speaking, *optimization* means maximization under restrictions; the best-known example is linear programming. In this paper the term optimization is also used when restricitons are absent or ignored. There are many mathematical techniques for finding optimal values for the decision variables of nonlinear implicit functions (such functions may indeed be formulated by simulation models), possibly with stochastic noise; examples of such techniques are genetic algorithms, simulated annealing, and tabu search. However, this paper is limited to Response Surface Methodology (RSM). This methodology combines regression analysis and DOE (see §3) with a hill-climbing technique called steepest ascent. A good overview of RSM including references is Vollebregt (1996).

First four *general characteristics of* RSM are considered; then some details.

(i) RSM relies on first-order and second-order polynomial regression metamodels, now called *response surfaces* (see Equation 1 in §3.2).
(ii) It uses the *statistical designs* of DOE (see §3.3).
(iii) It is augmented with the mathematical (not statistical) technique of *steepest ascent*, to determine in which direction the decision variables should be changed.

(iv) It uses the mathematical technique of *canonical analysis* to analyze the shape of the optimal region: does that region have a unique maximum, a saddle point or a ridge?

Next consider some details. RSM begins by selecting a *starting point.* Because RSM is a heuristic (no success guaranteed!), several starting points may be tried later on, if time permits.

RSM explores the *neighborhood* of that starting point. Locally the response surface is approximated by a first-order polynomial in the decision variables (Taylor series expansion).

The main effects β_h (see Equation 1) are estimated, using a design with $n \approx k+1$ (see §3.3.1). Suppose $\hat{\beta}_1 >> \hat{\beta}_2 > 0$. Then obviously the increase of decision variable 1 (say) z_I should be larger than that of z_2. The *steepest ascent path* means $\Delta z_1/\Delta z_2 = \hat{\beta}_1/\hat{\beta}_2$ (no standardization; also see next paragraph).

Unfortunately, the steepest ascent technique does not quantify the *step size* along this path. Therefore the analysts may try a specific value for the step size. If that value yields a lower response, then this value should be reduced. Otherwise, one more step is taken. Ultimately, the response must decrease, since the first-order polynomial is only an approximation. Then the procedure is *repeated*: around the best point so far, a new first-order polynomial is estimated, after simulating $n \approx k+1$ combinations of z_1 through z_k. And so on.

In the neighborhood of the top, a hyperplane can *not* be an adequate representation. To detect this lack of fit, the analysts may use cross-validation (see §3.2). Other diagnostic measures are the well-known measure R^2 and modern measures such as PRESS, discussed in Kleijnen (1987). So when a hyperplane no longer approximates the local I/O transformation well enough, then a second-order polynomial is fitted (see §3.3.4).

Finally, the *optimal* values of z_h are found by straightforward differentiation of the fitted quadratic polynomial. A more sophisticated evaluation is *canonical analysis.*

Consider the following *case study.* A decision support system (DSS) for production planning in a steel tube factory is simulated and is to be optimized. There are fourteen decision variables, and two response variables (namely, a production and a commercial criterion); one response variable is maximized, whereas the other one forms a side-restriction.

Simulation of one combination takes six hours of computer time, so searching for the optimal combination can not be performed using only common sense. Details can be found in Kleijnen (1993).

More applications can be found in Hood and Welch (1993), Kleijnen(1987, 1995d), and Kleijnen and Van Groenendaal (1992).

7. CONCLUSIONS

The problem addressed in this paper (see especially §1) is that there are *five related analyses*: (i) sensitivity analysis (SA) or what-if analysis, (ii) uncertainty analysis (UA) or risk analysis, (iii) screening, (iv) validation, and (v) optimization. And interesting questions are: *when* should *which* type of analysis be applied; which *statistical techniques* should then be used?

This paper gave a *survey* of these issues, emphasizing *statistical* procedures. Such procedures yield reproducible, objective, quantitative results.

Briefly, *sensitivity analysis* determines which model inputs are really important. From the users' perspective, the important inputs are either controllable or not. The controllable inputs may be optimized. The values of the uncontrollable inputs may be well-known, in which case these values can be used for validation of the model. If the values of the uncontrollable inputs are not well-known, then the likelihood of their values can be quantified objectively or subjectively, and the probability of specific output values can be quantified by *uncertainty* analysis.

More specifically, SA means that the model is subjected to *extreme value* testing. A model is valid only within its experimental frame (which was defined as the limited set of circumstances under which the real system is to be observed or experimented with). Mathematically that frame might be defined as the hypercube formed by the k standardized inputs $x_{i,h}$ of the model (more complicated definitions allow for restrictions such as linear conditions on the input combinations (say) $\sum_{h=1}^{k} x_{i,h} = 1$). Experimental designs such as 2^{k-p} fractional factorials specify *which* combinations are actually observed or simulated; for example, a 2^{-p} fraction of the 2^k corner points of that hypercube (although he did not discuss DOE, Draper 1995, p. 55 does speak of 'staking out the corners in model space'). The n observed input combinations and their corresponding responses are analyzed through a

regression (meta)model, which is an approximation of the simulation model's I/O transformation. That regression model quantifies the importance of the simulation inputs.

This paper proposed the following *five stages* in the analysis of a simulation model.

Stage 1. Obviously, *validation* is one of the first problems that must be solved in a simulation study. The availability of data on the real system determines the type of statistical technique to be used for validation. Regression analysis, however, may be applied, whether data are available or not, albeit through different regression models. See §4.

Stage 2. When the simulation study is still in its *pilot* phase, then very many inputs may be conceivably important. The really important inputs can be identified through Bettonvil and Kleijnen (1995)'s *sequential bifurcation*, which is a *screening* technique that is based on aggregation and sequential experimentation. See §2.

Stage 3. The important inputs found in stage 2 are investigated in a *more detailed sensitivity analysis*, including *interactions* and possibly *quadratic* effects (curvature). This investigation may use *design of experiments* (DOE), which includes classical designs such as 2^{k-p} fractional factorial designs and possibly central composite designs. Such designs give estimators of the effects in the *regression metamodel* that are better: minimum variance, unbiased linear estimators. See §3.

Stage 4. The important inputs should be split into two groups: inputs that are under the decision makers' control versus environmental inputs. Though the important *environmental inputs* cannot be controlled, information on the values they are likely to assume might be gathered. If the value of such an input is not precisely known, then the chances of various values can be quantified through a *probability function*. If a sample of data is available, then this function can be estimated objectively, applying mathematical statistics; otherwise subjective expert opinions are used. UA quantifies the uncertainties of the model outputs that result from the uncertainties in the model inputs. Output uncertainty is quantified through a statistical distribution. This analysis uses the Monte Carlo

technique. This Monte Carlo experiment has smaller variance when applying *Latin hypercube sampling* (LHS). This paper claims (controversially?) that LHS is a variance reduction technique (VRT), not a screening technique.

SA does not tell *how likely* a particular combination of inputs (specified by a statistical design) is, whereas UA does account for the probabilities of input values.

Note that Bayesians average the outcomes, using the probabilities of the various input scenarios; see Draper (1995). It might be argued, however, that in general it is the job of 'managers' (including governments) to use their 'intuition' to make decisions; it is the job of scientists to prepare a basis for such decisions.

Combining UA with regression analysis shows which non-controllable inputs contribute most to the uncertainty in the output.

UA of stochastic simulation models is an interesting area for further research. See §5.

Stage 5. The *policy variables* should be controlled. *Response Surface Methodology* (RSM) is a heuristic technique that combines DOE, regression analysis, and steepest ascent, in order to find the model inputs that give better model responses, possibly the best response. See §6.

Applications of the recommended techniques for these five stages are quite plentiful (see references).

An important conclusion is that SA should precede UA. Each type of analysis may apply its own set of statistical techniques, for example, 2^{k-p} fractional designs in SA, and LHS in UA. Some techniques may be applied in both analyses, for example, regression modeling. Hopefully, this paper succeeded in explaining *when* to use *which* technique! Yet, sensitivity and risk analyses remain *controversial* topics; communication within and among scientific disciplines is certainly needed.

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